

DENSE RANDOM GRAPHS: SPECTRA OF GRAPHONS

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NET WORKS

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Spectra of random matrices have been analysed for almost a century. In recent years, many interesting results have been derived for spectra of random matrices associated with networks.

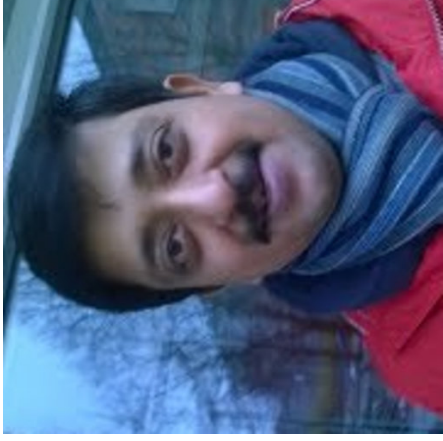
QUESTION ADDRESSED IN THIS TALK:

What can be said about the spectrum of the adjacency matrix of a large inhomogeneous Erdős-Rényi random graph?

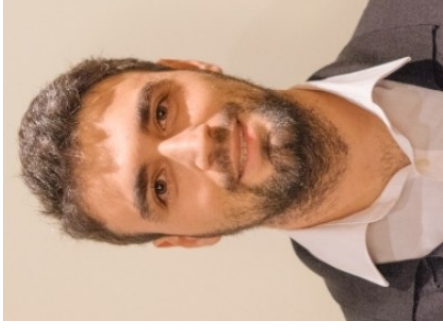




Arijit Chakrabarty



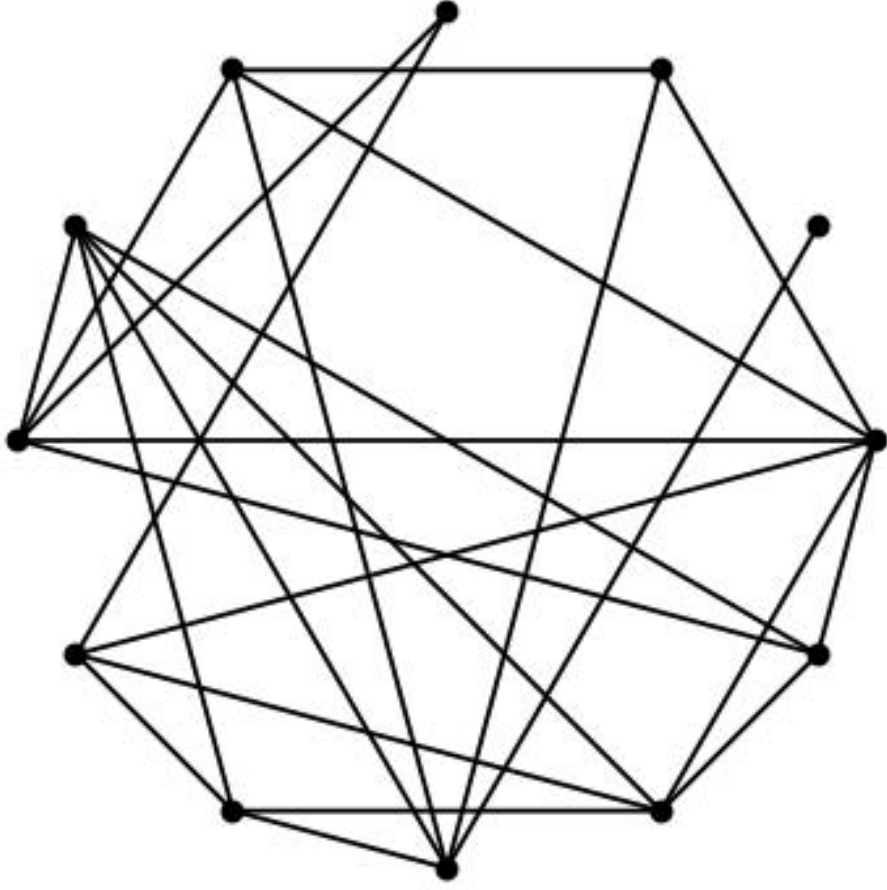
Rajat Hazra



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Random Matrices: Theory and Applications 2020

+ work in progress



Erdős-Rényi random graph: $n = 12, p = 0.3$

§ SETTING

1. Let $(\varepsilon_n)_{n \in \mathbb{N}}$ be a sequence of positive numbers such that

$$\lim_{n \rightarrow \infty} \varepsilon_n = 0, \quad \lim_{n \rightarrow \infty} n\varepsilon_n = \infty.$$

Let $r: [0, 1] \times [0, 1] \rightarrow [0, \infty)$ be a continuous function such that $r(x, y) = r(y, x)$ for all $x, y \in [0, 1]$.

2. Fix $n \in \mathbb{N}$, and consider the **inhomogeneous Erdős-Rényi random graph** ER_n on n vertices where an edge is placed between the pair of vertices $\{i, j\}$ with probability

$$\varepsilon_n r\left(\frac{i}{n}, \frac{j}{n}\right), \quad 1 \leq i, j \leq n,$$

independently for different edges. Write \mathbb{P} for the law of ER_n .

3. Let A_n be the adjacency matrix of ER_n . Write

$$\lambda_i(A_n), \quad 1 \leq i \leq n,$$

for the real eigenvalues of A_n . Define the empirical spectral distribution of A_n as

$$\text{ESD}(A_n) = \frac{1}{n} \sum_{i=1}^n \delta_{\lambda_i(A_n)},$$

which is a random probability distribution on \mathbb{R} .



§ SCALING

THEOREM 1:

There exists a *compactly supported symmetric probability measure* μ on \mathbb{R} such that, *weakly in \mathbb{P} -probability*,

$$\lim_{n \rightarrow \infty} \text{ESD} \left(A_n / \sqrt{n\epsilon_n} \right) = \mu.$$

Furthermore, if

$$\min_{x,y \in [0,1]} r(x,y) > 0,$$

then μ is *absolutely continuous* with respect to Lebesgue measure. The density of μ can be characterised *implicitly* via an integral equation for its Stieltjes transform.

It is possible to identify μ when r is of rank 1, i.e.,

$$r(x, y) = \nu(x)\nu(y), \quad x, y \in [0, 1],$$

for some continuous function $\nu: [0, 1] \rightarrow [0, \infty)$.

THEOREM 2:

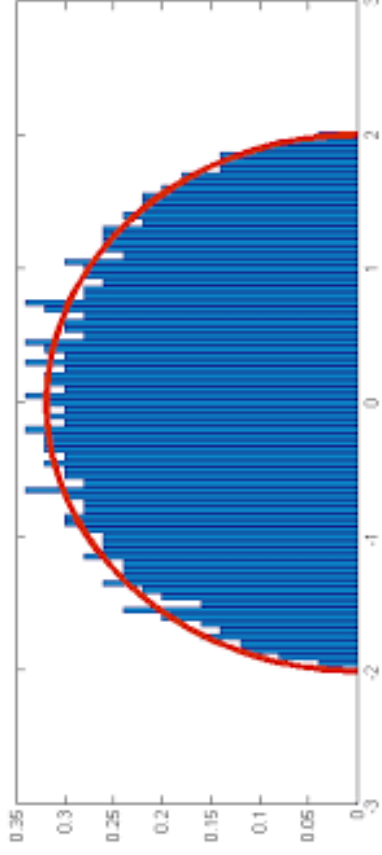
If r is of product form, then

$$\mu = \mu_\nu \boxtimes \mu_*$$

where

$$\begin{aligned} \mu_\nu &= \text{LAW}[\nu(U)], \quad U = \text{UNIF}[0, 1], \\ \mu_* &= \text{standard Wigner semicircle law,} \end{aligned}$$

and \boxtimes denotes *free multiplicative convolution*.



Wigner 1955

In free probability, the **Wigner semicircle law** takes over the role of the **normal law** in classical probability. The so-called **free cumulants** replace the classical cumulants, in the sense that partitions are replaced by **non-crossing partitions**.

Just as the cumulants of degree ≥ 2 are all zero if and only if the distribution is **normal**, the **free cumulants** of degree ≥ 2 are all zero if and only if the distribution is the **Wigner semicircle law**.

THEOREM 3:

Theorems 1–2 can be generalized to the situation where the function r is *random*, depends on n and converges to a deterministic limit as $n \rightarrow \infty$.

Key ingredients of the proof are:

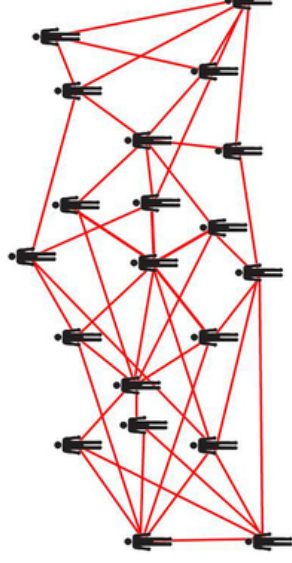
centering, Gaussianisation, perturbation, decoupling, combinatorics from free probability, ...



§ APPLICATION 1 statistics for social networks

Consider a community of n individuals, represented by the vertices in ER_n . Data is available about which individuals are acquainted. Based on this data, the sociability pattern of the community has to be inferred statistically.

Let ρ denote a probability measure on $[0, \infty)$ with bounded support. Let $(R_i)_{1 \leq i \leq n}$ be i.i.d. random variables drawn from ρ . Think of R_i as the sociability index of individual i .



Pick n so large that

$$0 \leq \varepsilon_n R_i R_j \leq 1 \quad \forall 1 \leq i, j \leq n.$$

Suppose that i, j are acquainted with probability $\varepsilon_n R_i R_j$, which is represented by an edge in ER_n between vertices i, j . The data that is available is the adjacency matrix A_n .

The statistical inference problem is to estimate ρ from A_n . To standardise ρ , we assume that w.o.l.g.

$$\int_0^\infty x \rho(dx) = 1.$$

Since, weakly \mathbb{P} -a.s.,

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \delta_{R_i} = \rho,$$

Theorem 3 gives that, weakly in \mathbb{P} -probability,

$$\lim_{n \rightarrow \infty} \text{ESD} \left(A_n / \sqrt{n\varepsilon_n} \right) = \rho \boxtimes \mu_*.$$

In practice, ε_n is unknown, which can be worked around by arguing that, weakly in \mathbb{P} -probability,

$$\lim_{n \rightarrow \infty} \text{ESD} \left(\sqrt{\frac{n}{\text{Tr}(A_n^2)}} A_n \right) = \rho \boxtimes \mu_*,$$

because, in \mathbb{P} -probability,

$$\text{Tr}(A_n^2) = \sum_{1 \leq i, j \leq n} A_n(i, j) A_n(j, i) = \sum_{1 \leq i, j \leq n} A_n(i, j)^2 \sim n^2 \varepsilon_n.$$

The procedure is that $\rho \boxtimes \mu_*$ can be statistically estimated from A_n . Subsequently, ρ can be estimated because the moments of $\rho \boxtimes \mu_*$ are functions of the moments of ρ and μ_* .

Indeed, since the moments of μ_* are known, the moments of ρ can be recursively computed from the moments of $\rho \boxtimes \mu_*$. Since ρ is compactly supported, it can in turn be computed via its moments.



§ APPLICATION 2 configuration model

Let \mathcal{G}_n be the set of simple graphs on n vertices. We fix the degrees of all the vertices, namely, vertex i has degree d_i^* , where

$$\vec{d}_n^* = \{d_i^*\}_{1 \leq i \leq n}$$

is a sequence of positive integers of which we only require that it is graphical, i.e., there is at least one simple graph matching these degrees.



hard configuration model

The soft ensemble P_n is the unique probability distribution on \mathcal{G}_n with the following two properties:

(I) The average degree of vertex i , defined by

$$\sum_{G \in \mathcal{G}_n} d_i(G) P_n(G),$$

equals d_i^* for all $i \leq n$.

(II) The entropy of P_n , defined by

$$- \sum_{G \in \mathcal{G}_n} P_n(G) \log P_n(G),$$

is maximal.

P_n models a random graph of which no prior information is available other than the average degrees.

soft configuration model

Property (II) forces P_n to take the form Jaynes 1957

$$P_n(G) = \frac{1}{Z_n(\vec{\theta}^*)} \exp \left[- \sum_{i=1}^n \theta_i^* d_i(G) \right], \quad G \in \mathcal{G}_n,$$

where $\vec{\theta}_n^* = \{\theta_i^*\}_{1 \leq i \leq n}$ is the unique sequence of Lagrange multipliers such that property (I) is satisfied.

Reparametrisation yields

$$P_n(G) = \prod_{1 \leq i < j \leq n} (p_{ij}^*)^{A_n[G](i,j)} (1 - p_{ij}^*)^{1 - A_n[G](i,j)}, \quad G \in \mathcal{G}_n,$$

where $A_n[G]$ is the adjacency matrix of G , and

$$p_{ij}^* = \frac{x_i^* x_j^*}{1 + x_i^* x_j^*}, \quad x_i^* = e^{-\theta_i^*}, \quad 1 \leq i \neq j \leq n.$$



Property (I) requires that

$$d_i^* = \sum_{\substack{1 \leq j \leq n \\ j \neq i}} p_{ij}^*, \quad 1 \leq i \leq n,$$

which constitutes a set of n equations for n unknowns.

Abbreviate

$$m_n = \max_{1 \leq i \leq n} d_i^*.$$

We focus on the regime

$$\lim_{n \rightarrow \infty} m_n = \infty, \quad \lim_{n \rightarrow \infty} m_n / \sqrt{n} = 0.$$

It turns out that in this regime

$$p_{ij}^* = [1 + o(1)] \frac{d_i^* d_j^*}{\sigma_n}, \quad n \rightarrow \infty,$$

with

$$\sigma_n = \sum_{1 \leq i \leq n} d_i^*.$$

Pick

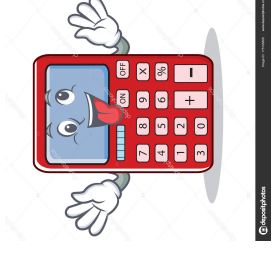
$$\varepsilon_n = m_n^2 / \sigma_n.$$

Then

$$\lim_{n \rightarrow \infty} \varepsilon_n = 0, \quad \lim_{n \rightarrow \infty} n \varepsilon_n = \infty,$$

and

$$p_{ij}^* = [1 + o(1)] \varepsilon_n (d_i^* / m_n) (d_j^* / m_n).$$



Under the assumption that

$$\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n \delta_{d_i^*} / m_n = \rho$$

for some probability measure ρ , Theorem 3 gives that, weakly in \mathbb{P} -probability,

$$\lim_{n \rightarrow \infty} \text{ESD}(A_n / \sqrt{n \varepsilon_n}) = \rho \boxtimes \mu_*$$

This identifies the scaling of the ESD for the network that is modeled by the soft configuration model as a function of the imposed average degrees.

§ LARGEST EIGENVALUE

Suppose that $\epsilon_n \equiv 1$ and that the reference graphon r satisfies

$$r \in (0, 1) \text{ is continuous a.e.,} \\ \log r, \log(1 - r) \in L^1([0, 1]^2).$$

Then we can use the LDP for the inhomogeneous Erdős-Rényi random graph.

Let $\lambda_1 = \lambda_1(\tilde{h}^G)$ denote the maximal eigenvalue of the adjacency matrix $A_n[G]$ of G .

LDP

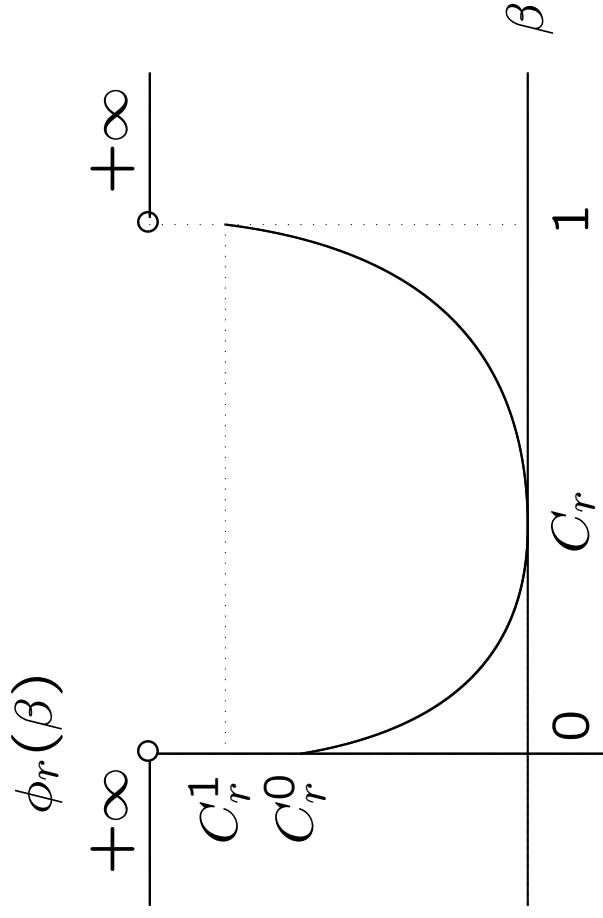
THEOREM 4: Chakrabarty, Hazra, dH & Sfragara in progress

The family $(\mathbb{P}_{n,r}(\lambda_1/n \in \cdot))_{n \in \mathbb{N}}$ satisfies the LDP on \mathbb{R} with rate $\binom{n}{2}$ and with rate function ϕ_r given by

$$\phi_r(\beta) = \inf_{\substack{h \in \mathcal{W} \\ \|h\| = \beta}} I_r(h)$$

with $\|h\|$ the operator norm of h .

The claim is an immediate consequence of the LDP and the contraction principle, because $\tilde{h}^G \mapsto \lambda_1(\tilde{h}^G)/n$ is bounded and continuous with respect to the δ_{\square} -metric.



$$C_r = \|r\|$$

$$C_r^0 = \int_{[0,1]^2} dx dy \log \frac{1}{1-r(x,y)}$$

$$C_r^1 = \int_{[0,1]^2} dx dy \log \frac{1}{r(x,y)}$$

HEURISTICS:

When $\beta = C_r$, the graphon h that minimizes $I_r(h)$ such that $\|h\| = C_r$ is the reference graphon $h = r$. When $\beta > C_r$, we need a graphon h with a larger operator norm. The large deviation cannot go above 1, which is represented by the constant graphon $h \equiv 1$, for which $I_r(1) = C_r^1$. A similar observation holds when $\beta < C_r$.

THEOREM 5: Chakrabarty, Hazra, dH & Sfragara in progress

- (i) ϕ_r is continuous on $[0, 1]$.
- (ii) ϕ_r is strictly unimodal on $[0, 1]$.



If the reference graphon r is of rank 1, i.e.,

$$r(x, y) = \nu(x) \nu(y) \quad \text{for some } \nu: [0, 1] \rightarrow [0, 1],$$

then more can be said. Define

$$m_k = \int_{[0,1]} dx \nu(x)^k, \quad k \in \mathbb{N}.$$

The following theorem relies on a closer analysis of the variational formula for ϕ_r .

THEOREM 6: Chakrabarty, Hazra, dH & Sfragara in progress

(i) For $\beta \rightarrow C_r$,

$$\phi_r(\beta) \sim K_r (\beta - C_r)^2$$

with

$$C_r = m_1^2, \quad K_r = \frac{1}{2} \frac{m_1^4}{m_3^2 - m_4^2}.$$

(ii) For $\beta \uparrow 1$,

$$C_r^1 - \phi_r(\beta) \sim (1 - \beta) \log \frac{1}{1 - \beta}.$$

(iii) For $\beta \downarrow 0$,

$$C_r^0 - \phi_r(\beta) \sim \beta \log \frac{1}{\beta}.$$

§ EXPANSION

A key tool in the proof of the last theorems is a **series expansion** for the operator norm of a graphon h around any graphon of rank 1.

THEOREM 7: Chakrabarty, Hazra, dH & Sfragata in progress

Let $\bar{h} \in \mathcal{W}$ be of the form $\bar{h}(x, y) = \bar{v}(x)\bar{v}(y)$, $x, y \in [0, 1]$. For any $h \in \mathcal{W}$ such that $\|h - \bar{h}\| < \|h\|$, the operator norm $\|h\|$ is the unique solution of the equation

$$\|h\| = \sum_{n \in \mathbb{N}_0} \frac{1}{\|h\|^n} \mathcal{F}_n(h, \bar{h}),$$

where

$$\mathcal{F}_n(h, \bar{h}) = \int_{[0,1]^2} dx dy \bar{v}(x)(h - \bar{h})^{\otimes n}(x, y)\bar{v}(y)$$

with \otimes denoting convolution.

PROOF



We have

$$hu = \lambda u,$$

where λ is the norm and the maximal eigenvalue of h , and u is an eigenfunction of h . Put $g = h - \bar{h}$, and write

$$\begin{aligned}(g + \bar{h})u = \lambda u &\longrightarrow (\lambda - g)u = \bar{h}u \\ \longrightarrow u = (\lambda - g)^{-1}\bar{h}u &\longrightarrow u = (\lambda - g)^{-1}\bar{v}\langle\bar{v}, u\rangle.\end{aligned}$$

Hence

$$\begin{aligned}\langle\bar{v}, u\rangle = \langle\bar{v}, u\rangle\langle\bar{v}, (\lambda - g)^{-1}\bar{v}\rangle &\longrightarrow 1 = \langle\bar{v}, (\lambda - g)^{-1}\bar{v}\rangle \\ \longrightarrow \lambda = \langle\bar{v}, (1 - g/\lambda)^{-1}\bar{v}\rangle,\end{aligned}$$

where $\lambda - g$ is invertible because $\|g\| = \|h - \bar{h}\| < \|h\| = \lambda$, and we take $u > 0$ so that $\langle \bar{v}, u \rangle \neq 0$. We can expand

$$\begin{aligned}
 \lambda &= \left\langle \bar{v}, \sum_{n \in \mathbb{N}_0} (g/\lambda)^n \bar{v} \right\rangle \\
 &= \sum_{n \in \mathbb{N}_0} \frac{1}{\lambda^n} \int_{[0,1]^{n+1}} dx_0 dx_1 \cdots dx_n \\
 &\quad \bar{v}(x_0) g(x_0, x_1) \times \cdots \times g(x_{n-1}, x_n) \bar{v}(x_n) \\
 &= \sum_{n \in \mathbb{N}_0} \frac{1}{\lambda^n} \mathcal{F}_n(h, \bar{h}).
 \end{aligned}$$

Since $\|h\| = \lambda$, we get the claim. \square

