

Ordinary random walk

Perturbed random walk

Moments of first passage times and related quantities for ordinary and perturbed random walks

Alexander Iksanov, Kiev, Ukraine 💻

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Gerold Alsmeyer (Münster)

[1] Alsmeyer, G., Iksanov, A. and Meiners, M. (2014). Power and exponential moments of the number of visits and related quantities for perturbed random walks. *JTP*, to appear.

[2] Iksanov, A. and Meiners, M. (2010). Exponential moments of first passage times and related quantities for random walks. *ECP.* **15**, 365–375.

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Matthias Meiners (Darmstadt)

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- definition of the main objects
- simple cases and previously known results
- criteria for the finiteness of exponential moments



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# Ordinary random walk: definition of the main objects

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Let  $(\xi_k)_{k\in\mathbb{N}}$  be independent copies of a real-valued random variable  $\xi$ .

A random sequence  $(S_n)_{n \in \mathbb{N}_0}$  defined by

 $S_0 := 0, \ S_n := \xi_1 + \ldots + \xi_n, \ n \in \mathbb{N}$ 

is called zero-delayed (ordinary) random walk.



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### Main objects

• the first passage time into  $(x, \infty)$ 

 $\tau(x) := \inf\{n \in \mathbb{N}_0 : S_n > x\}, x \ge 0;$ 

• the *number of visits* to the interval  $(-\infty, x]$ 

$$N(x) := \#\{n \in \mathbb{N} : S_n \le x\} \\ = \sum_{n \ge 1} \mathbb{1}_{\{S_n \le x\}}, \quad x \ge 0;$$

• the *last exit time* from  $(-\infty, x]$ 

$$\rho(x) = \sup\{n \in \mathbb{N}_0 : S_n \le x\}, \quad x \ge 0.$$

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When  $\mathbb{P}\{\xi \ge 0\} = 1$  and  $\mathbb{P}\{\xi = 0\} < 1$ ,  $\tau(x) - 1 = N(x) = \rho(x), \quad x \ge 0.$ 

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When  $\mathbb{P}\{\xi \ge 0\} = 1$  and  $\mathbb{P}\{\xi = 0\} < 1$ ,  $\tau(x) - 1 = N(x) = \rho(x), \quad x > 0.$ 

**Proposition** (Beljaev & Maksimov (1963)) Assume that  $\mathbb{P}\{\xi \ge 0\} = 1$  and let  $\beta := \mathbb{P}\{\xi = 0\} \in [0, 1)$ . Then for a > 0 the following conditions are equivalent:

 $\mathbb{E}e^{a\tau(x)} < \infty$  for some (hence every)  $x \ge 0$ ;

$$a < -\log\beta$$

where  $-\log \beta := \infty$  if  $\beta = 0$ . The same holds for N(x) and  $\rho(x)$ .



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where  $-\log \beta := \infty$  if  $\beta = 0$ . The same holds for N(x) and  $\rho(x)$ . **Proof**: If  $a < -\log \beta$ , pick r > 0 s.t.  $e^a \mathbb{E} e^{-r\xi} < 1$  $\sum_{k \ge 0} e^{ak} \mathbb{P}\{\tau(x) > k\} = \sum_{k \ge 0} e^{ak} \mathbb{P}\{S_k \le x\}$   $\leq e^{rx} \sum_{k \ge 0} (e^a \mathbb{E} e^{-r\xi})^k < \infty$ 



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$$\begin{array}{rcl} \text{If } a \geq -\log \beta, \\ & \sum_{k \geq 0} e^{ak} \mathbb{P}\{S_k \leq x\} & \geq & \sum_{k \geq 0} e^{ak} \mathbb{P}\{S_k = 0\} \\ & = & \sum_{k \geq 0} e^{ak} \beta^k = \infty. \end{array}$$

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General case:  $\mathbb{P}\{\xi < 0\} \ge 0$ . *FINITENESS:* • Borovkov (1962): if

$$a < R := -\log \inf_{t \ge 0} \mathbb{E}e^{-t\xi},$$

then 
$$\mathbb{E}e^{a\tau(0)} < \infty$$
;  
if  $a > R$ , then  $\mathbb{E}e^{a\tau(0)} = \infty$ .



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then  $\mathbb{E}e^{a\tau(0)} < \infty$ ; if a > R, then  $\mathbb{E}e^{a\tau(0)} = \infty$ . • Heyde (1964): assuming that  $\mathbb{E}\xi \in (0,\infty)$ ,  $\mathbb{E}e^{a\tau(x)} < \infty$  for some a > 0 iff

$$\mathbb{E}e^{b\xi^-} < \infty$$
 for some  $b > 0$ .



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• Doney (1989); Bertoin & Doney (1996) – other relevant results.



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• Doney (1989); Bertoin & Doney (1996) – other relevant results.

ASYMPTOTICS:

• Heyde (1966): whenever  $\mathbb{E}\xi \in (0,\infty)$  and  $\mathbb{E}e^{a\tau(0)} < \infty$ ,

 $\mathbb{E}e^{a\tau(x)} \sim \operatorname{const} e^{(a/\mathbb{E}\xi)x}, \ x \to \infty.$ 

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Theorem (I. & Meiners (2010)) Let a > 0 and  $\mathbb{P}\{\xi < 0\} > 0$ . Then the following conditions are equivalent:

$$\begin{split} \mathbb{E}e^{a\tau(x)} &< \infty \text{ for some (hence every) } x \geq 0; \\ \mathbb{E}e^{aN(x)} &< \infty \quad \text{for some (every)} \quad x \geq 0; \\ \sum_{n\geq 1} \frac{e^{an}}{n} \mathbb{P}\{S_n \leq x\} &< \infty \quad \text{for some (every)} \quad x \geq 0; \end{split}$$

$$a \le R := -\log \inf_{t \ge 0} \mathbb{E}e^{-t\xi}$$

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$$\begin{split} \mathbb{E}e^{a\rho(x)} &< \infty \quad \text{for some (hence every)} \quad x \ge 0; \\ \sum_{n \ge 0} e^{an} \mathbb{P}\{S_n \le x\} < \infty \quad \text{for some (every)} \quad x \ge 0; \\ \text{either } a &< R = -\log \inf_{t \ge 0} \mathbb{E}e^{-t\xi}, \text{ or} \\ a &= R \text{ and } \mathbb{E}\xi e^{-\gamma_0 \xi} > 0, \end{split}$$

where  $\gamma_0$  is the unique positive number such that  $\mathbb{E}e^{-\gamma_0\xi} = e^{-R}$ .



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**Example** (simple random walk) Let  $p \in (1/2, 1)$  and

 $\mathbb{P}\{\xi = 1\} = p = 1 - \mathbb{P}\{\xi = -1\} =: 1 - q.$ 

#### Then

$$\begin{split} \mathbb{E}e^{-t\xi} &= pe^{-t} + qe^t \text{ and } \\ R &= -\log\inf_{t\geq 0} \mathbb{E}e^{-t\xi} = -\log(2\sqrt{pq}). \end{split}$$

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 $\mathbb{P}\{\xi = 1\} = p = 1 - \mathbb{P}\{\xi = -1\} =: 1 - q.$ 

Then

$$\mathbb{E}e^{-t\xi} = pe^{-t} + qe^t \text{ and}$$
$$R = -\log \inf_{t \ge 0} \mathbb{E}e^{-t\xi} = -\log(2\sqrt{pq}).$$

(2m)

One can check that

$$\mathbb{P}\{\tau(0) = 2n - 1\} = \frac{1}{2q} \frac{\binom{2n}{n}}{2^{2n}(2n - 1)} (2\sqrt{pq})^{2n}$$
$$= \frac{1}{2q} \frac{\binom{2n}{n}}{2^{2n}(2n - 1)} e^{-2nR}, \quad n \in \mathbb{N};$$
$$\mathbb{P}\{\rho(0) = 2n\} = (p - q) \binom{2n}{n} (pq)^{n}$$
$$= (p - q) \frac{\binom{2n}{n}}{2^{2n}} e^{-2nR}, \quad n \in \mathbb{N}_{0}.$$

Stirling's formula yields

$$\frac{\binom{2n}{n}}{2^{2n}} \sim \frac{1}{\sqrt{\pi n}}, \ n \to \infty,$$

which implies that

$$\mathbb{E}e^{R au(0)} < \infty$$
 and  $\mathbb{E}e^{R
ho(0)} = \infty$ 

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NONNEGATIVE CASE

Theorem (I. & Meiners (2010)) Let a > 0,  $\mathbb{P}\{\xi \ge 0\} = 1$  and  $\mathbb{P}\{\xi = 0\} < 1$ . Assume that  $\mathbb{E}e^{a\tau(x)} < \infty$  for some (hence every)  $x \ge 0$ . Then, as  $x \to \infty$ ,  $\mathbb{E}e^{a\tau(x)} \sim \operatorname{const} e^{\gamma x}$ 

where

 $\mathrm{const} \; = \; \begin{cases} \frac{1-e^{-a}}{\gamma\mathbb{E}\xi e^{-\gamma\xi}}, & \text{if } \mathcal{L}(\xi) \text{ is non-lattice,} \\ \frac{\lambda(1-e^{-a})}{(1-e^{-\lambda\gamma})\mathbb{E}\xi e^{-\gamma\xi}}, & \text{if } \mathcal{L}(\xi) \text{ is } \lambda\text{-lattice} \end{cases}$ 

 $\boldsymbol{\gamma}$  is the unique positive number such that

$$\mathbb{E}e^{-\gamma\xi} = e^{-a},$$

and in the  $\lambda$ -lattice case the limit is taken over  $x \in \lambda \mathbb{N}$ .



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GENERAL CASE

#### Theorem (I. & Meiners(2010))

Let a > 0 and  $\mathbb{P}\{\xi < 0\} > 0$ . Assume that  $\mathbb{E}e^{a\rho(x)} < \infty$  for some (hence every)  $x \ge 0$ . Then, as  $x \to \infty$ ,  $\mathbb{E}e^{a\rho(x)} \sim \operatorname{const} e^{\gamma x}$ 

where

# $\mathrm{const} \ = \ \begin{cases} \frac{e^{-a}(1-\mathbb{E}e^{-\gamma M^+})}{\gamma\mathbb{E}\xi e^{-\gamma\xi}}, & \text{if } \mathcal{L}(\xi) \text{ is non-lattice,} \\ \frac{\lambda e^{-a}(1-\mathbb{E}e^{-\gamma M^+})}{(1-e^{-\lambda\gamma})\mathbb{E}\xi e^{-\gamma\xi}}, & \text{if } \mathcal{L}(\xi) \text{ is } \lambda\text{-lattice} \end{cases}$

 $M:=\inf_{n\,>\,1}S_n;\,\gamma$  is a minimal positive number such that

 $\mathbb{E}e^{-\gamma\xi} = e^{-a}.$ 

In the  $\lambda$ -lattice case the limit is taken over  $x \in \lambda \mathbb{N}$ .

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where

$$\mathrm{const} \; = \; \begin{cases} \frac{e^{-a}(1 - \mathbb{E}e^{-\gamma M}^+)}{\gamma \mathbb{E}\xi e^{-\gamma \xi}}, & \text{if } \mathcal{L}(\xi) \text{ is non-lattice} \\ \frac{\lambda e^{-a}(1 - \mathbb{E}e^{-\gamma M}^+)}{(1 - e^{-\lambda \gamma}) \mathbb{E}\xi e^{-\gamma \xi}}, & \text{if } \mathcal{L}(\xi) \text{ is } \lambda \text{-lattice} \end{cases}$$

 $M:=\inf_{n\geq 1}S_n;\,\gamma$  is a minimal positive number such that

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### FINITENESS

```
(Kesten & Maller (1996)) For p > 0
\mathbb{E}(\tau(x))^{p+1} < \infty \quad \Leftrightarrow \quad \mathbb{E}(N(x))^p < \infty \quad \Leftrightarrow\mathbb{E}(\rho(x))^p < \infty.
```

(!) It may happen that there exists q > 0

$$\lim_{x\to\infty} x^q \mathbb{P}\{\tau>x\} = 0 \ \text{ and } \ \liminf_{x\to\infty} x^q \mathbb{P}\{N>x\} > 0.$$



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(!) It may happen that there exists q > 0

 $\lim_{x\to\infty} x^q \mathbb{P}\{\tau>x\}=0 \ \, \text{and} \ \, \liminf_{x\to\infty} x^q \mathbb{P}\{N>x\}>0.$ 

(I. & Meiners (2010)) Provided that

$$R = -\log \inf_{t \ge 0} \mathbb{E}e^{-t\xi} > 0,$$

it holds that

$$\begin{split} \mathbb{E} e^{a\tau(x)} < \infty, \ \mathbb{E} e^{aN(x)} < \infty \text{ iff } a \leq R, \text{ and} \\ \mathbb{E} e^{a\rho(x)} < \infty \text{ if } a < R, \end{split}$$

whereas  $\mathbb{E}e^{R\rho(x)}$  is finite in some cases and infinite in others.

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Asymptotics

Let T(x) denote any of the rvs  $\tau(x), \ N(x)$  or  $\rho(x).$ 

(Kesten & Maller (1996)) If  $\mathbb{E}(T(0))^p < \infty$  then

$$\mathbb{E}(T(x))^p \asymp \left(\frac{x}{\mathbb{E}\min(\xi^+, x)}\right)^p, x \to \infty.$$



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- Illustrating example
- Asymptotics of the exponential moments
- Exponential moments vs power moments

Perturbed random walk

Asymptotics

Let T(x) denote any of the rvs  $\tau(x), \ N(x)$  or  $\rho(x).$ 

(Kesten & Maller (1996)) If  $\mathbb{E}(T(0))^p < \infty$  then

$$\mathbb{E}(T(x))^p \asymp \left(\frac{x}{\mathbb{E}\min(\xi^+, x)}\right)^p, x \to \infty.$$

(I. & Meiners (2010)) If  $\mathbb{E}e^{aT(0)} < \infty$  then

$$\mathbb{E}e^{aT(x)} \simeq e^{\gamma x}, \quad x \to \infty.$$

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Criteria for the finiteness of moments of the number of visits

Random processes with immigration Let  $(\xi_k, \eta_k)_{k \in \mathbb{N}}$  be independent copies of a random vector  $(\xi, \eta)$  with arbitrary dependence between  $\xi$  and  $\eta$ .

By a perturbed random walk is meant a random sequence  $(T_n)_{n\in\mathbb{N}}$  defined by

 $T_n := S_{n-1} + \eta_n = \xi_1 + \ldots + \xi_{n-1} + \eta_n, \quad n \in \mathbb{N}.$ 



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The so defined PRW is related to

- perpetuities
- shot noise processes
- the Bernoulli sieve
- GI/G/∞ queues



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The so defined PRW is related to

perpetuities

$$\sum_{n\geq 1} e^{T_n}$$

- shot noise processes
- the Bernoulli sieve
- $GI/G/\infty$  queues



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Random processes with immigration

$$T_n = S_{n-1} + \eta_n, \quad n \in \mathbb{N}.$$

**Proposition** (Goldie & Maller (2000)) Either  $\lim_{n\to\infty} T_n = +\infty$  a.s. (*drift to*  $+\infty$ ), or  $\lim_{n\to\infty} T_n = -\infty$  a.s. (*drift to*  $-\infty$ ), or  $-\infty = \liminf_{n\to\infty} T_n < \limsup_{n\to\infty} T_n = +\infty$  a.s. (*oscillation*).

(!) It may happen that the perturbed random walk  $T_n$  oscillates when the corresponding ordinary random walk  $S_n$  drifts to  $\pm\infty$ .



### Perturbed random walk: Criteria for the finiteness of moments of the last exit time

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Random processes with immigration  $T_n = S_{n-1} + \eta_n, \quad n \in \mathbb{N}.$ 

For  $x \in \mathbb{R}$ , the *last exit time* from  $(-\infty, x]$  is

$$\rho^*(x) = \begin{cases} \sup\{n \in \mathbb{N} : T_n \le x\}, & \text{if } \inf_{k \ge 1} T_k \le x, \\ 0, & \text{if } \inf_{k \ge 1} T_k > x \end{cases}$$



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For  $x \in \mathbb{R}$ , the *last exit time* from  $(-\infty, x]$  is

$$p^*(x) = \begin{cases} \sup\{n \in \mathbb{N} : T_n \le x\}, & \text{if } \inf_{k \ge 1} T_k \le x, \\ 0, & \text{if } \inf_{k \ge 1} T_k > x \end{cases}$$

Power moments.

**Theorem (Alsmeyer, I. & Meiners (2014))** Let  $(T_n)$  be positively divergent and p > 0. The following assertions are equivalent:

$$\begin{split} \mathbb{E}(\rho^*(x))^p &< \infty \text{ for some (hence every) } x \in \mathbb{R};\\ \mathbb{E}(\rho(y))^p &< \infty \quad \text{for some (hence every) } y \geq 0 \quad \text{and}\\ \mathbb{E}J^{p+1}(\eta^-) &< \infty,\\ \text{where } J(z) &:= \frac{z}{\mathbb{E}(\xi^+ \wedge z)}, \ z > 0, \text{ and}\\ \rho(y) &= \sup\{n \in \mathbb{N}_0 : S_n \leq y\}, \ y \geq 0. \end{split}$$



### Perturbed random walk: Criteria for the finiteness of moments of the last exit time

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EXPONENTIAL MOMENTS.

Theorem (Alsmeyer, I. & Meiners (2014)) Let  $(T_n)$  be positively divergent,  $\mathbb{P}\{\xi < 0\} > 0$  and a > 0. The following assertions are equivalent:

$$\begin{split} \mathbb{E} \exp(a\rho^*(x)) &< \infty \text{ for some (hence every) } x \in \mathbb{R}; \\ \mathbb{E} \exp(a\rho(y)) &< \infty \text{ for some (hence every) } y \geq 0 \quad \text{and} \\ \mathbb{E} e^{-\gamma\eta} &< \infty, \end{split}$$

where  $\gamma$  is the minimal positive number satisfying  $\mathbb{E} e^{-\gamma\xi}=e^{-a},$  and

$$\rho(y) = \sup\{n \in \mathbb{N}_0 : S_n \le y\}, \quad y \ge 0.$$



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Random processes with immigration Alsmeyer, I. & Meiners (2014): for a > 0

$$\begin{split} \mathbb{E} e^{a\rho^*(x)} &< \infty \quad \text{for some (hence every}) x \in \mathbb{R} \quad \Leftrightarrow \\ \sum_{n \geq 1} e^{an} \mathbb{P} \{ T_n \leq x \} &< \infty \quad \text{for some/all } x \in \mathbb{R}. \end{split}$$



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Kesten & Maller (1996): for p > 0

$$\begin{split} \mathbb{E}\rho(x)^p &< \infty \quad \text{for some (hence every) } x \geq 0 \quad \Leftrightarrow \\ \sum_{n\geq 1} n^{p-1} \mathbb{P}\{S_n \leq x\} &< \infty \quad \text{for some/all } x \geq 0. \end{split}$$



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However

$$\mathbb{E}\rho^*(x)^p < \infty \quad \Leftrightarrow \quad \sum_{n \ge 1} n^{p-1} \mathbb{P}\{T_n \le x\} < \infty !!!$$



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Random processes with immigration However

$$\mathbb{E}\rho^*(x)^p < \infty \quad \Leftrightarrow \quad \sum_{n>1} n^{p-1} \mathbb{P}\{T_n \le x\} < \infty !!!$$

Indeed,

 $\sum_{n\geq 1} n^{p-1} \mathbb{P}\{T_n\leq x\}<\infty \quad \Leftrightarrow \quad \mathbb{E}\rho(x)^p<\infty \ \text{ and } \ \mathbb{E}J(\eta^-)^p<\infty,$ 

whereas

 $\mathbb{E}\rho^*(x)^p < \infty \quad \Leftrightarrow \quad \mathbb{E}\rho(x)^p < \infty \quad \text{and} \quad \mathbb{E}J(\eta^-)^{p+1} < \infty.$ 



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 $T_n = S_{n-1} + \eta_n, \quad n \in \mathbb{N}.$ 

For  $x \in \mathbb{R}$ , the *number of visits* to  $(-\infty, x]$  is

$$N^*(x) = \#\{n \in \mathbb{N} : T_n \le x\}$$

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Power moments.

**Theorem** (Alsmeyer, I. & Meiners (2014)) Let  $(T_n)$  be positively divergent and p > 0. The following assertions are equivalent:

 $\mathbb{E}(N^*(x))^p < \infty$  for some (hence every)  $x \in \mathbb{R}$ ;

$$\mathbb{E}(N(y))^p < \infty$$
 for some (hence every)  $y \ge 0$ .

Here

$$N(y) = \#\{n \in \mathbb{N}_0 : S_n \le y\}, y \ge 0.$$

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EXPONENTIAL MOMENTS.

**Theorem** (Alsmeyer, I. & Meiners (2014)) Let  $(T_n)$  be positively divergent.

(a) If  $\mathbb{P}\{\xi < 0\} > 0$ , then the following assertions are equivalent:

 $\mathbb{E} \exp(aN^*(x)) < \infty$  for some (hence every)  $x \in \mathbb{R}$ ,  $\mathbb{E} \exp(aN(y)) < \infty$  for some (hence every) $y \ge 0$ .

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EXPONENTIAL MOMENTS.

**Theorem** (Alsmeyer, I. & Meiners (2014)) Let  $(T_n)$  be positively divergent.

(b) If  $\xi \ge 0$  a.s., then the assertions

 $\mathbb{E}\exp(aN^*(x)) < \infty,$  $e^a \mathbb{P}\{\xi = 0, \eta \le x\} + \mathbb{P}\{\xi = 0, \eta > x\} < 1$ 

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are equivalent for each a > 0 and  $x \in \mathbb{R}$ .

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DEFINITION:

•  $X = (X(t))_{t \in \mathbb{R}}$  is a random process with paths in the Skorokhod space  $D(\mathbb{R})$ ;



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- $(S_n)_{n\in\mathbb{N}_0}$  is the zero-delayed random walk with jumps  $\xi_k$ , i.e.,

$$S_0 := 0, \ S_n := \xi_1 + \xi_2 + \ldots + \xi_n, \ n \in \mathbb{N}.$$



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$$S_0 := 0, \ S_n := \xi_1 + \xi_2 + \ldots + \xi_n, \ n \in \mathbb{N}.$$

### Random process with immigration

The process  $Y := (Y(t))_{t \in \mathbb{R}}$  defined by

$$Y(t) := \sum_{k \ge 0} X_{k+1}(t - S_k), \ t \in \mathbb{R}$$

will be called random process with immigration at the epochs of a renewal process.

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INTERPRETATION in the case when  $\xi > 0$  a.s. and X(t) = 0, t < 0.



At time  $S_0 = 0$  the immigrant 1 arrives at the system and runs a random process  $X_1$ .





INTERPRETATION in the case when  $\xi > 0$  a.s. and X(t) = 0, t < 0.



At time  $S_0 = 0$  the immigrant 1 arrives at the system and runs a random process  $X_1$ .

At time  $S_k$  the immigrant k + 1 arrives at the system and runs a random process  $X_{k+1}$  which evolves along with already existing processes  $X_1, \ldots, X_k$ .

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INTERPRETATION: in the case when  $\xi > 0$  a.s. and X(t) = 0, t < 0.



Y(t) is then the sum of all processes run by the immigrants up to and including time  $t, \mbox{ i.e.},$ 

$$Y(t) = \sum_{k \ge 0} X_{k+1}(t - S_k).$$



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$$N^{*}(t) = \#\{k \in \mathbb{N} : T_{k} \leq t\}$$
  
=  $\sum_{k \geq 0} \mathbb{1}_{\{\eta_{k+1} \leq t - S_{k}\}}$   
=  $\sum_{k \geq 0} X_{k+1}(t - S_{k}) = Y(t), t \in \mathbb{R}.$ 

Therefore  $(N^*(t))_{t \in \mathbb{R}}$  is a random process with immigration with  $X_k(t) = \mathbb{1}_{\{\eta_k \leq t\}}, k \in \mathbb{N}$ .

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### THANK YOU FOR YOUR ATTENTION

Alexander Iksanov